# The Monte Carlo method Introduction

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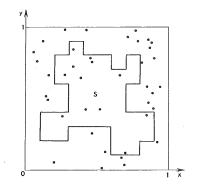
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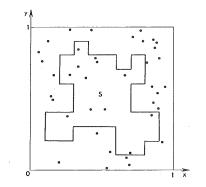
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It is obvious that the area of S should be equal approximately to the ratio N'/N.

In our example N=40 of which N'=12 is inside S, thus

$$\frac{N'}{N} = \frac{12}{40} = 0.30,$$

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In practise the Monte Carlo (MC) method is not used to calculate areas of flat figures, as there are much better methods, though a bit more complicated, which are much better in this respect. However, the MC method becomes indispensable, if we want to calculate *volume* of a body in a multidimensional space.

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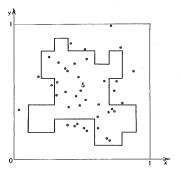
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#### How the random points should actually be selected?

Consider our example once more, but this time imagine that we fix our figure on the wall and let some dart player to hit it with darts from some distance.

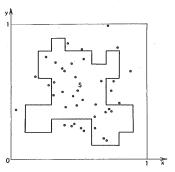


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#### Random variables

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A random variable  $\xi$  which takes on values from a discrete set  $\{x_1, x_2, ..., x_n\}$  is specified by the table

$$\xi = \left(\begin{array}{cccc} x_1 & x_2 & \dots & x_n \\ p_1 & p_2 & \dots & p_n \end{array}\right),$$

where  $p_1, p_2, ..., p_n$  are the corresponding probabilities.

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- ① all  $p_i$ 's are non negative:  $p_i > 0$ ;
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If c is an arbitrary non-random number, then

$$M(\xi + c) = M\xi + c,$$
  
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If  $\xi$  and  $\eta$  are two arbitrary random variables, then

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If in the course of many observations of the random variable  $\xi$  we obtain a set of values  $\xi_1, \xi_2, ..., \xi_N$ , each of them being equal to one of the numbers  $x_1, x_2, ..., x_n$ , then the arithmetic mean of these numbers will be close to  $M\xi$ , i.e.

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Using the definition and the above properties of the expected value variance  $D\xi$  can be written in the following way

$$D\xi = M[(\xi - M\xi)^2] = M[\xi^2 - 2M\xi \cdot \xi + (M\xi)^2] = M(\xi^2) - 2M\xi \cdot M\xi + (M\xi)^2],$$

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### Basic properties of variance

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# Examples

**Example.** Let us consider a random variable  $\kappa$  which is the number of points obtained in a single throw of a die. It can be represented by the following table

$$\kappa = \left(\begin{array}{ccccc} 1 & 2 & 3 & 4 & 5 & 6 \\ \frac{1}{6} & \frac{1}{6} & \frac{1}{6} & \frac{1}{6} & \frac{1}{6} & \frac{1}{6} \end{array}\right).$$

Let us find the expected value and variance of  $\kappa$ .

$$\mathbf{M}\kappa = \sum_{i=1}^{6} x_i p_i = 1 \cdot \frac{1}{6} + 2 \cdot \frac{1}{6} + \dots + 6 \cdot \frac{1}{6} = 3.5,$$

$$\mathbf{D}\kappa = \mathbf{M}(\kappa^2) - (\mathbf{M}\kappa)^2 = 1^2 \cdot \frac{1}{6} + 2^2 \cdot \frac{1}{6} + \dots + 6^2 \cdot \frac{1}{6} - 3.5^2 = 2.917.$$

# **Examples**

**Example.** Let us consider a random variable  $\theta$  which is a single toss of a coin, assuming that a toss of heads brings 3 points an that of tails brings 4 points. It can be represented by the following table

$$\theta = \left(\begin{array}{cc} 3 & 4 \\ \frac{1}{2} & \frac{1}{2} \end{array}\right).$$

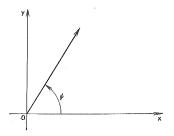
The expected value and variance of  $\theta$  are

$$\mathbf{M}\theta = \sum_{i=1}^{2} x_{i} p_{i} = 3 \cdot \frac{1}{2} + 4 \cdot \frac{1}{2} = 3.5,$$

$$\mathbf{D}\theta = \mathbf{M}(\theta^{2}) - (\mathbf{M}\theta)^{2} = 3^{2} \cdot \frac{1}{2} + 4^{2} \cdot \frac{1}{2} - 3.5^{2} = 0.25.$$

We see that  $\pmb{M}\theta = \pmb{M}\kappa$ , but  $\pmb{D}\theta < \pmb{D}\kappa$ . The inequality can be easily understood, as the maximum deviation of  $\theta$  from the mean 3.5 is  $\pm 0.5$ , while the maximum deviation of  $\kappa$  from the mean is  $\pm 2.5$ .

Let us assume that a certain amount of radium is placed on a plane at the origin of the coordinates. An  $\alpha$ -particle is emitted in each decay of a radium atom. The direction of motion of this particle will be characterized by an angle  $\psi$ , see the figure below.



Since any direction of emission is possible, the random variable  $\psi$  can assume any value fro 0 to  $2\pi$ .

We shall call a random variable  $\xi$  continuous if it takes on any value out of a certain interval (a, b).

A continuous random variable  $\xi$  is defined by specifying the interval (a,b) of its variation and the function p(x) called the *probability density* of the random variable  $\xi$ .

The physical meaning of p(x) is the following. Let (a',b') be an arbitrary interval contained within (a,b), i.e.  $a \le a'$  and  $b' \le b$ . Then the probability that  $\xi$  falls inside (a',b') is given by the integral

$$\mathbf{P}\{a'<\xi< b'\}=\int\limits_{a'}^{b'}p(x)dx.$$

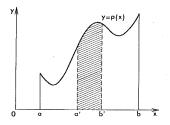
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$$\mathbf{P}\{a'<\xi< b'\}=\int_{a'}^{b'}p(x)dx.$$

The integral  $\int_{a'}^{b'} p(x)dx$  is equal to the hatched area in the following figure.



The set of values of  $\xi$  can be any interval. The cases when  $a=-\infty$  as well as  $b=+\infty$  are also possible.

However, the density p(x) must satisfy two conditions, similar to those of probabilities  $p_i$  in the case of a discrete random variable:

- the density p(x) is positive: p(x) > 0,
- 2 the integral of the density p(x) over the whole interval (a, b) is equal to 1:

$$\int_{a}^{b} p(x)dx = 1.$$

The value

$$\mathbf{M}\xi = \int_{a}^{b} x p(x) dx$$

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All properties of the expected value and variance of the continuous random variable are the same as those of the discrete random variable.

We will mention here only one more expression for the expected value of a function of  $\xi$ . Let us assume that the random variable  $\xi$  is characterized by the probability density p(x) and consider an arbitrary continuous function f(x). We introduce a random variable  $\eta = f(\xi)$ . It can be shown that

$$\mathbf{M}f(\xi) = \int_{a}^{b} f(x)p(x)dx.$$

We must emphasize that in general  $Mf(\xi) \neq f(M\xi)$ .

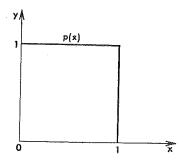
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A random variable  $\gamma$ , defined on the interval (0,1) with the density p(x) = 1 is said to be *uniformly distributed* in (0,1), see the figure below.



Indeed, for any interval (a',b') within (0,1) the probability that  $\gamma$  will take on the value within (a',b') is equal to

$$\int_{a'}^{b'} p(x)dx = b' - a',$$

i.e. to the length of this interval. If, in particular, we divide the interval (0,1) into an arbitrary number of intervals of equal length, the probability of  $\gamma$  falling within one of these intervals will be the same. Note that

$$\mathbf{M}\gamma = \int_{0}^{1} xp(x)dx = \int_{0}^{1} xdx = \frac{1}{2},$$

$$\mathbf{D}\gamma = \int_{0}^{1} x^{2}p(x)dx - (\mathbf{M}\gamma)^{2} = \frac{1}{3} - \frac{1}{4} = \frac{1}{12}.$$

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$$\int_{a'}^{b'} p(x) dx = b' - a',$$

i.e. to the length of this interval. If, in particular, we divide the interval (0,1) into an arbitrary number of intervals of equal length, the probability of  $\gamma$  falling within one of these intervals will be the same. Note that

$$\mathbf{M}\gamma = \int_{0}^{1} xp(x)dx = \int_{0}^{1} xdx = \frac{1}{2},$$

$$\mathbf{D}\gamma = \int_{0}^{1} x^{2}p(x)dx - (\mathbf{M}\gamma)^{2} = \frac{1}{3} - \frac{1}{4} = \frac{1}{12}.$$